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Extent of Surgical Resection in Lower-Grade Gliomas: Differential Impact Based on Molecular Subtype

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ABSTRACT

BACKGROUND AND PURPOSE: Diffuse lower-grade gliomas are classified into prognostically meaningful molecular subtypes. We aimed to determine the impact of surgical resection on overall survival in lower-grade glioma molecular subtypes.

MATERIALS AND METHODS: For 172 patients with lower-grade gliomas (World Health Organization grade II or III), pre- and postsurgical glioma volumes were determined using a semiautomated segmentation software based on FLAIR or T2-weighted MR imaging sequences. The association of pre- and postsurgical glioma volume and the percentage of glioma resection with overall survival was determined for the entire cohort and separately for lower-grade glioma molecular subtypes based on *isocitrate dehydrogenase (IDH)* and 1p/19q status, after adjustment for age, sex, World Health Organization grade, chemotherapy administration, and radiation therapy administration.

RESULTS: For the entire cohort, postsurgical glioma volume (hazard ratio, 1.80; 95% CI, 1.18–2.75; $P = .006$) and the percentage of resection (hazard ratio, 3.22; 95% CI, 1.79–5.82; $P < .001$) were associated with overall survival. For *IDH*-mutant 1p/19q-codeleted oligodendrogliomas, the percentage of resection (hazard ratio, 6.69; 95% CI, 1.57–28.46; $P = .01$) was associated with overall survival. For *IDH*-mutant 1p/19q-noncodeleted astrocytomas, presurgical glioma volume (hazard ratio, 3.20; 95% CI, 1.22–8.39; $P = .018$), postsurgical glioma volume (hazard ratio, 2.33; 95% CI, 1.32–4.12; $P = .004$), and percentage of resection (hazard ratio, 4.34; 95% CI, 1.74–10.81; $P = .002$) were associated with overall survival. For *IDH*-wild-type lower-grade gliomas, pre-/postsurgical glioma volume and percentage of resection were not associated with overall survival.

CONCLUSIONS: The extent of surgical resection has a differential survival impact in patients with lower-grade gliomas based on their molecular subtype. *IDH*-mutant lower-grade gliomas benefit from a greater extent of surgical resection, with the strongest impact observed for *IDH*-mutant 1p/19q-noncodeleted astrocytomas.

ABBREVIATIONS: LGG = lower-grade glioma; mut = mutation; WHO = World Health Organization; *IDHmut-Noncodeled* = LGGs with an *IDH* mutation but lacking 1p/19q codeletion; *IDHmut-Codel* = LGGs with an *IDH* mutation and codeletion of chromosome arms 1p and 19q; *IDHwt* = LGGs in *IDH*-wild-type subgroup

Diffuse lower-grade gliomas (LGGs) are World Health Organization (WHO) grade II and III infiltrative brain neoplasms. In 2016, the WHO updated the classification of diffuse LGGs by integrating molecular markers with histopathology.¹ Between 70% and 80% of LGGs are now known to have a prognostically favorable mutation in the *isocitrate dehydrogenase (IDH)* gene.^{2,3} Oligodendrogliomas (grade II) and anaplastic oligoden-

drogliomas (grade III) are defined by the presence of both an *IDH* mutation and codeletion of chromosome arms 1p and 19q (heretofore collectively referred to as *IDHmut-Codel*). LGGs with an *IDH* mutation but lacking 1p/19q codeletion include diffuse astrocytomas (grade II) and anaplastic astrocytomas (grade III) (heretofore collectively referred to as *IDHmut-Noncodeled*).^{1,3–5} These *IDH*-mutant LGG subgroups have distinct prognostic and therapeutic profiles.^{3–7} The *IDH*-wild-type subgroup of LGGs (heretofore collectively referred to as *IDHwt*) is associated with the most aggressive clinical behavior and worst outcome, similar to that of glioblastomas (WHO grade IV), though heterogeneity in the behavior and molecular features within this subgroup is being increasingly recognized.^{8–11}


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Multiple studies have provided evidence supporting the maximum safe surgical resection as the front-line treatment of diffuse LGGs.^{12–19} However, most studies were undertaken before the 2016 WHO update for the classification of diffuse LGGs, without accounting for defining molecular markers. It is unknown to what extent these earlier study results are confounded by the use of older nonmolecular classification schema. Only recently have volumetric MR imaging studies incorporated *IDH* and 1p/19q-codeletion status when assessing the impact of surgical resection on patient outcomes in LGGs.^{20,21}

The purpose of our study was to determine the effect of surgical resection extent on overall survival in a patient cohort with LGGs stratified by molecular subtype. Among the *IDH*-mutant LGGs, we hypothesized that the extent of surgery might have greater impact among *IDHmut-Noncodel* subtypes than the *IDHmut-Codel* subtypes, given the overall worse prognosis among the former subtype. We hypothesized that surgery would have prognostic impact among the *IDHwt* subtypes as well, though these results might be confounded by the molecular heterogeneity of this subtype.

MATERIALS AND METHODS

This retrospective study of patient data was Health Insurance Portability and Accountability Act–compliant and was approved by our institutional review board (University of Virginia Health System).

Patient Selection

The patients included in this study were selected from an institutional neuro-oncology/neuroradiology diffuse glioma data base maintained at our institution, containing a total of 429 diffuse glioma cases diagnosed between 2000 and 2018. All patients who are seen by Medical Neuro-Oncology at our institution are added to our data base. From this data base, patients were selected on the basis of the following inclusion criteria: 1) diagnosed with a diffuse LGG (WHO grades II and III), 2) a known molecular subtype based on *IDH* and 1p/19q-codeletion status, 3) available presurgical and postsurgical MR images, and 4) known chemotherapy and radiation therapy history at the last follow-up. After review of the electronic medical record and PACS, we excluded 176 cases with grade IV histology (ie, glioblastomas) and 49 LGGs with unknown molecular status. Of the remaining 204 cases, we excluded 20 cases for lack of pre- or postsurgical MR imaging and 12 cases for unknown chemotherapy or radiation therapy history at time of the last follow-up. A total of 172 patients met the criteria for study inclusion. The mean time interval between the presurgical MR imaging and the operation was 5.18 days, and the mean time interval between the postsurgical MR imaging and the operation was 79.3 days. As per the methodology of Wijnenga et al,²⁰ we preferentially avoided using the immediate postsurgical MR imaging scans for performing volume measurements to avoid including postsurgical edema or ischemia in our measurements.

Neuroimaging Analysis

Pre- and postsurgical MR images were analyzed by a board-certified neuroradiologist with a Certificate of Added Qualification in diagnostic neuroradiology and 6 years of experience. Measure-

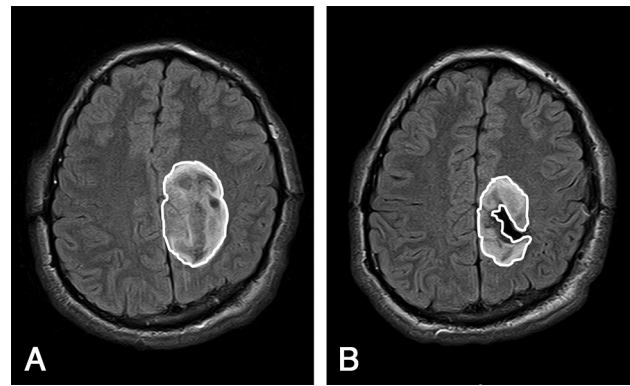


FIG 1. Glioma segmentation shown on axial FLAIR images of pre-surgical MR imaging (A) and postsurgical MR imaging (B) in a patient with a diffuse *IDH*-mutant astrocytoma who underwent subtotal resection.

ments of glioma volume on the presurgical and postsurgical MR imaging examinations were undertaken using the semiautomated lesion-management tool in the PACS using FLAIR or T2-weighted imaging (3D sequences were used when possible). MR imaging examinations analyzed in this study were performed on various 1.5T and 3T MR imaging scanners, and pulse sequences with variable parameters were used. On 3T MR imaging, the 3D-T2WI and FLAIR sequences used the following parameters—3D-T2WI: FOV = 256 mm, slice thickness = 1.0 mm, matrix = 256 × 256, TR = 3200 ms, TE = 413 ms, NEX = 1; 3D-FLAIR: FOV = 256 mm, slice thickness = 1.0 mm, matrix = 256 × 256, TR = 5000 ms, TI = 1800 ms, TE = 386 ms, NEX = 1. To create glioma volumes of interest, the neuroradiologist reader manually traced the contour of the glioma using axial images from the highest quality FLAIR or T2-weighted sequence on the presurgical and postsurgical MR images (Fig 1), and the software calculated a 3D volume based on the neuroradiologist's segmentation. For the presurgical volume measurements, in cases in which peritumoral edema could not be confidently distinguished from infiltrative glioma, the neuroradiologist reader erred on the side of including any high signal present on the T2/FLAIR sequences within the volume measurements. Based on presurgical and postsurgical glioma volumes, the percentage of glioma resection was calculated as $100\% \times (1 - [\text{Postsurgical Glioma Volume}] / [\text{Presurgical Glioma Volume}])$.

Neuropathology

Glioma histology and grade and *IDH* and 1p/19q statuses were retrieved from the electronic medical record. *IDH* and 1p/19q statuses were tested in the Clinical Laboratory Improvement Amendments–certified molecular pathology laboratory at our institution. *IDH*-mutation status was first tested by immunohistochemistry using an *IDH1* R132H-mutant-specific antibody. Immunohistochemistry was performed on 4-μm-thick sections from formalin-fixed paraffin-embedded tissue following the manufacturer's recommended protocol (Bond-III; Leica Biosystem, Newcastle Upon Tyne, UK). Commercially purchased antibodies against *IDH1* (R132H) mutant protein (DIA-H09; Dianova, Hamburg, Germany) were used.^{22,23} In cases with negative findings on immunohistochemistry, *IDH1/2*

Table 1: Summary of patient characteristics^a

	<i>IDHmut-Codel</i> (n = 68)	<i>IDHmut-Noncodel</i> (n = 63)	<i>IDHwt</i> (n = 41)
Sex			
Female	37 (54.4%)	30 (47.6%)	18 (43.9%)
Male	31 (45.6%)	33 (52.4%)	23 (56.1%)
Age (yr) ^a	45.5 (35.5–53)	33 (27–42.8)	57.0 (49.5–63.3)
WHO grade			
II	50 (73.5%)	46 (73.0%)	25 (61.0%)
III	18 (26.5%)	17 (27.0%)	16 (39.0%)
Chemotherapy			
Yes	55 (80.9%)	53 (84.1%)	30 (73.2%)
No	13 (19.1%)	10 (15.9%)	11 (26.8%)
Radiation therapy			
Yes	45 (66.2%)	53 (84.1%)	31 (75.6%)
No	23 (33.8%)	10 (15.9%)	10 (24.4%)
Presurgical glioma volume (cm ³) ^a	56.4 (26.5–111.6)	55.0 (31.0–116.2)	47.7 (18.0–86.4)
Postsurgical glioma volume (cm ³) ^a	28.4 (13.2–59.1)	29.3 (1.2–65.6)	30.1 (8.4–64.4)
Mortality events	12 (17.6%)	14 (22.2%)	13 (31.7%)

^a Data are listed as median and interquartile range. All other data are listed as absolute values and percentages.

mutation status was assessed by the clinically validated DNA pyrosequencing assay, using the PyroMark Q24 system, following the manufacturer's recommended protocol (QIAGEN; <https://www.qiagen.com/us/products/discovery-and-translational-research/pyrosequencing/instruments/pyromark-q24/#orderinginformation>).

We used the following primers—*IDH1* forward primer: 5'-Biot.-CATAATGTTGGCGTCAAAATGTG-3'; *IDH1* reverse primer: 5'-ACATGCAAAATCACATTATTGCC-3'; *IDH1* sequencing primer: 5'-TGATCCCCATAAGCAT-3'; *IDH2* forward primer: 5'-GTTCAAGCTGAAGAAGATGTGG-3'; *IDH2* reverse primer: 5'-Biot.-GTGGCCTTGACTGCAGAG-3'; *IDH2* sequencing primer: and 5'-AGCCCATCACATTGG-3'. The pyrosequencing assay is designed to detect mutations within codon 132 of *IDH1* and codon 172 of *IDH2*, as described previously.²⁴ The 1p/19q-codeletion status was determined using dual-color fluorescence in situ hybridization on formalin-fixed paraffin-embedded tissue. Commercial human probes were applied to localize 1p36, 1q25, 19p13 (Vysis; Abbott Molecular, Abbott Park, Illinois), and DAPI (Insitus Biotechnologies, Albuquerque, New Mexico) was used as a nuclear counterstain.

Statistical Analysis

Categoric variables are summarized by frequencies and percentages, and molecular subtype comparisons of categoric variables were conducted using the Fisher exact test. Continuous scale variables are summarized by the median, interquartile range, and range of the distribution, and molecular subtype comparisons of continuous scale variables were performed using the Wilcoxon rank sum test. Multivariate Cox proportional hazards regression was used to examine whether presurgical glioma volume, postsurgical glioma volume, and the percentage of glioma resection are uniquely associated with overall survival time for all cases and for cases in the 3 molecular subtypes. In total, 12 multivariate Cox models were constructed, and each model was identically specified with the only between-model difference being the variable of interest (presurgical glioma volume, postsurgical volume, or percentage of glioma resection) and the molecular subtype

(*IDHmut-Codel*, *IDHmut-Noncodel*, *IDHwt*, or all molecular subtypes combined). Per variable of interest (eg, presurgical glioma volume) and per molecular subtype of patients (eg, the *IDHmut-Noncodel* molecular subtype), the regression model included not only the variable of interest (eg, presurgical glioma volume) but also patient age and sex, WHO grades II and III, chemotherapy administration status (yes, no), and radiation therapy administration status (yes, no) as concomitant variable predictors of survival time. For all 12 multivariate Cox models, the survival times of patients who were known to be alive at last follow-up were treated as right-censored observations. With regard to hypothesis testing, the type III version of

the Wald χ^2 statistic served as the pivotal quantity for testing the null hypothesis that there is no unique association between the variable of interest (eg, presurgical glioma volume) and survival time after accounting for patient age and sex, WHO grade, chemotherapy administration status, and radiation therapy administration status associations with survival time. A $P \leq .05$ decision rule was established a priori as the null hypothesis rejection rule for testing for predictor variable-versus-survival time unique association, and the strength of the predictor variable versus survival time unique association was quantified by the adjusted hazard ratio.

RESULTS

Among the 172 study patients, there were 85 females (49.4%) and 87 males (50.6%). The median age was 43 years (interquartile range, 31–54 years; range, 17–76 years). Patient characteristics per molecular subtype are shown in Table 1. There were significant differences in age in each of the 3 molecular subgroups ($P \leq .001$), with patients with *IDHwt* being the oldest and those with *IDHmut-Noncodel* being the youngest. There were a total of 39 deaths. Median overall survival time was 16.5 years among patients with *IDHmut-Codel*, 12.1 years among those with *IDHmut-Noncodel*, and 6.3 years among those with *IDHwt*. Overall survival curves for the LGG subtypes are shown in Fig 2. Overall survival curves differed between patients with *IDHmut-Codel* and *IDHwt* ($P < .001$) and between those with *IDHmut-Noncodel* and *IDHwt* ($P < .001$). Overall survival curves did not significantly differ between patients with *IDHmut-Codel* and *IDHmut-Noncodel* ($P = .125$).

Among all cases, the median presurgical glioma volume was 54.7 cm³ (interquartile range, 27.6–107.6 cm³; range, 1.1–340 cm³) and the median postsurgical glioma volume was 29.7 cm³ (interquartile range, 6.5–62.8 cm³; range, 0–339 cm³). There was no significant difference in presurgical glioma volume or postsurgical glioma volume among the 3 LGG molecular subtypes. Patients with *IDHwt* were associated with a smaller percentage of resection compared with those with *IDHmut-Noncodel* ($P = .01$).

Associations among presurgical glioma volume, postsurgical

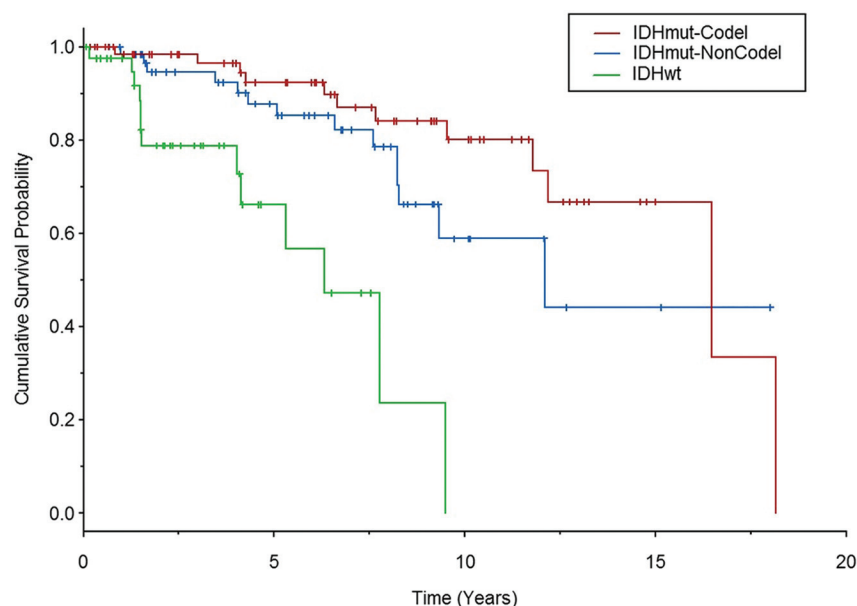


FIG 2. Kaplan-Meier curves for overall patient survival based on the LGG molecular subtype. Vertical line segments along the curve identify right-censored survival times. *IDHwt* LGGs were associated with significantly worse overall survival compared with both *IDH*-mutant subtypes.

Table 2: Survival analysis^a

Predictor	Ratio	Adjusted Hazard Ratio (95% CI)	P Value
All LGGs			
Presurgical glioma volume	3rd Q:1st Q	1.14 (0.79–1.64)	.479
Postsurgical glioma volume	3rd Q:1st Q	1.80 (1.18–2.75)	.006
Percentage glioma resection	25%:75%	3.22 (1.79–5.82)	<.001
<i>IDHmut-Codel</i>			
Presurgical glioma volume	3rd Q:1st Q	1.11 (0.53–2.33)	.781
Postsurgical glioma volume	3rd Q:1st Q	1.67 (0.82–3.48)	.170
Percentage glioma resection	25%:75%	6.69 (1.57–28.46)	.010
<i>IDHmut-Noncodel</i>			
Presurgical glioma volume	3rd Q:1st Q	3.20 (1.22–8.39)	.018
Postsurgical glioma volume	3rd Q:1st Q	2.33 (1.32–4.12)	.004
Percentage glioma resection	25%:75%	4.34 (1.74–10.81)	.002
<i>IDHwt</i>			
Presurgical glioma volume	3rd Q:1st Q	2.18 (0.47–10.15)	.319
Postsurgical glioma volume	3rd Q:1st Q	1.34 (0.28–4.76)	.653
Percentage glioma resection	25%:75%	0.91 (0.29–2.82)	.874

Note:—Q indicates quartile.

^a Predictor variables for overall survival are presurgical glioma volume, postsurgical glioma volume, and percentage glioma resection. Adjusted hazard ratios are listed for the entire cohort and for the 3 molecular subtypes, with age, sex, glioma grade, and chemotherapy and radiation therapy administration as the adjustment variables.

glioma volume, and the percentage of glioma resection and overall survival are shown in Table 2, after adjustment for patient age, sex, glioma grade, and chemo-radiation therapy administration. Among all LGGs in the cohort, postsurgical glioma volume ($P = .006$) and the percentage of glioma resection ($P < .001$) were associated with overall survival. However, associations varied among the 3 LGG molecular subtypes. For those with *IDHmut-Codel*, the percentage of resection was associated with overall survival ($P = .01$), but pre- or postsurgical glioma volumes were not. For those with *IDHmut-Noncodel*, presurgical glioma volume ($P = .018$), postsurgical glioma volume ($P = .004$), and the percentage of glioma resection ($P = .002$) were each associated with overall survival. For the *IDHwt* subtype, there was no association of presurgical glioma volume, postsurgical glioma volume, or

percentage of glioma resection with overall survival. Figure 3 shows overall survival curves according to the percentage of glioma resection for each of the molecular subtypes, and the On-line Figure shows overall survival curves according to postsurgical glioma volume strata. There were 13 patients with 0 mL of postsurgical residual glioma (*IDHmut-Codel*, $n = 1$; *IDHmut-Noncodel*, $n = 9$; *IDHwt*, $n = 3$), and all these patients were alive at the time of last follow-up.

DISCUSSION

Our study supports an association between a greater degree of surgical resection and overall survival in patients with *IDH*-mutant LGGs. We note an association between the extent of surgical resection and overall survival among both molecularly defined *IDH*-mutant astrocytomas and oligodendrogliomas; however, the impact of surgical resection extent on overall survival appears stronger for astrocytomas. The reason for the relatively stronger association with astrocytomas might relate to the relatively better responsiveness to nonsurgical therapies for oligodendrogliomas or oligodendrogliomas generally having longer survival times than astrocytomas, thus making it more difficult to demonstrate a survival benefit with an operation.^{3–7} We further note that the percentage of glioma resection was more strongly associated with overall survival than presurgical glioma volume or postsurgical glioma volume. Finally, we found no association between surgical resection and overall survival in *IDHwt* LGGs.

Prior studies have investigated the impact of surgical resection on the out-

comes of patients with LGGs. Among the older studies, many used nonvolumetric approximations of residual glioma extent.^{15,25–27} Additionally, most prior studies were undertaken before the 2016 WHO update for CNS neoplasm classification and thus did not strictly classify their patient cohorts by *IDH* and 1p/19q-codeletion status.^{13,14,16} Only recently have studies provided evidence of the impact of surgical resection after stratification by defining molecular markers.^{20,21,28,29}

The most important of this recent literature is the work from Wijnenga et al, in 2018,²⁰ who evaluated the impact of surgery in a large cohort of molecularly defined grade II diffuse gliomas, determining that larger postoperative glioma volume was associated with worse overall survival (hazard ratio, 1.01 per 1 cm³ increase in residual glioma volume). They noted that the amount

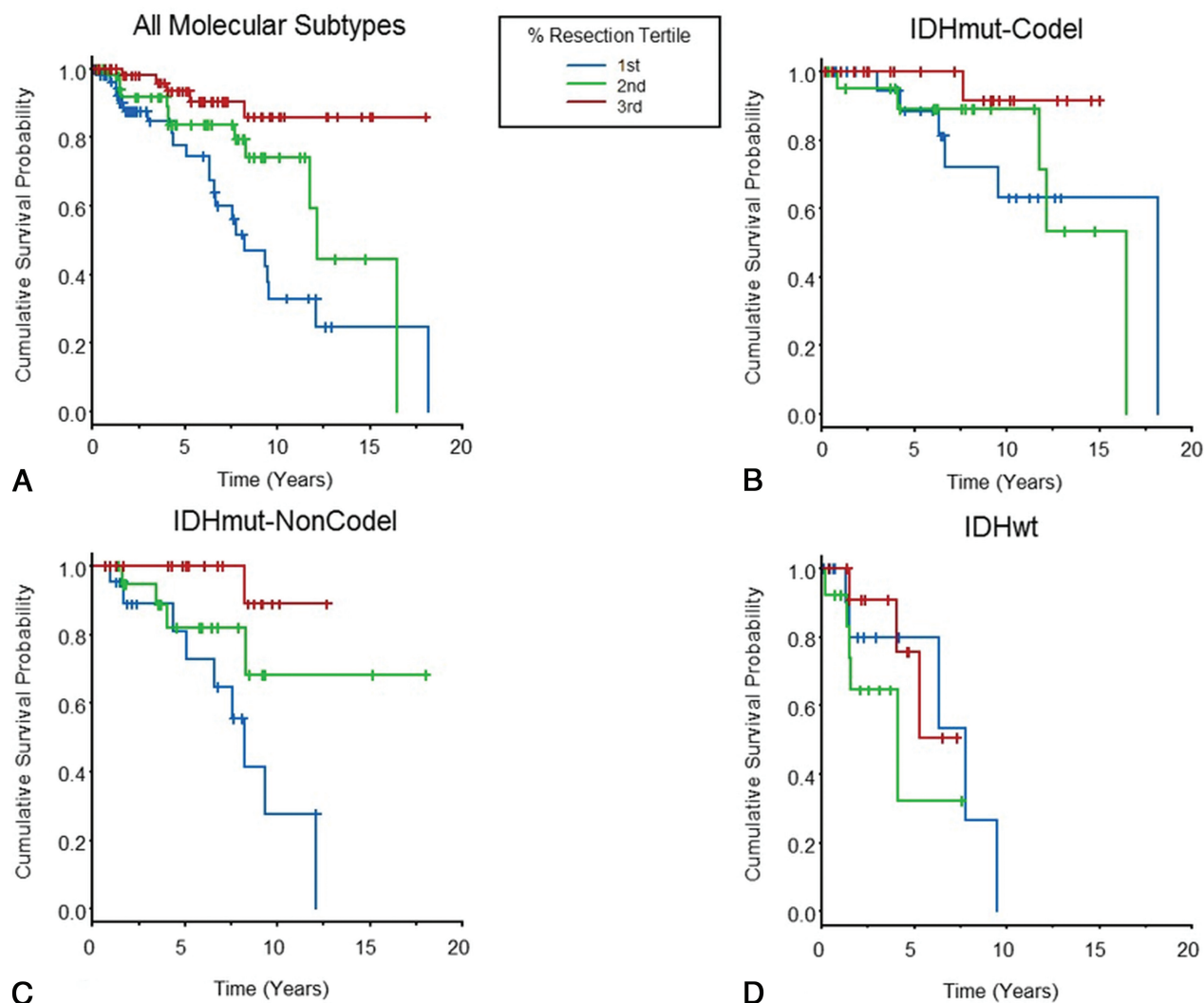


FIG 3. Kaplan-Meier curves in which overall patient survival is evaluated according to the percentage of glioma resection tertiles. Vertical line segments along the curve identify right-censored survival times. Data are shown for all patients with LGGs in the cohort (A) and separately for the 3 molecular subtypes (B–D).

of residual postsurgical glioma was more strongly associated with survival among *IDH*-mutant astrocytomas compared with oligodendrogliomas, similar to our results. Kawaguchi et al, in 2016,²⁸ evaluated a cohort of WHO grade III diffuse gliomas, undertaking an impressively comprehensive molecular analysis of their glioma specimens given the timing of their study. They found that gross total resection (as determined by postoperative MR imaging analysis) was significantly associated with improved overall survival among the *IDHmut-Noncodel* subgroup, but not for the *IDHmut-Codel* or the *IDHwt* subgroup. Study limitations included their binary stratification of surgical resection (yes/no gross total resection) and their choice of measuring residual glioma differently for enhancing and nonenhancing gliomas.

Eseonu et al, in 2017,²¹ evaluated a cohort of low-grade (presumably WHO grade II) diffuse gliomas. Molecular information was only available in a 51-patient subset of the patient cohort, among whom the extent of resection was associated with improved survival for the codeleted oligodendroglioma subgroup and *IDH*-wild-type subgroup (they did not have sufficient *IDH*-

mutant astrocytomas for analysis). Patel et al, in 2018,²⁹ evaluated a cohort of 74 patients with WHO grade II diffuse gliomas and, most interesting, found that the extent of glioma resection was associated with overall survival for the *IDH*-wild-type subgroup but not for the *IDH*-mutant subgroup, differing from our findings. Study limitations included lack of a description of *IDH*-mutation testing (with apparent lack of *IDH2*-mutation testing) and lack of stratification by 1p/19q-codeletion status.

The primary advantages of our study design include strict molecular classification of our patients with LGG, consistent with WHO guidelines, as well as volumetric measurement of pre- and postsurgical gliomas. Unlike most prior studies, we included both WHO grade II and III LGGs because WHO grade assignment among LGGs is subject to sampling bias and known to have significant interobserver variability.³⁰ Moreover, we ensured that our statistical analysis controlled for tumor grade. We also controlled for additional potentially confounding variables such as patient age, sex, and chemo-radiation therapy administration by the time of last follow-up. Our finding that the extent of surgical resection impacts overall survival in patients with *IDH*-mutant

LGGs, particularly in patients with *IDH*mut-Noncodel LGGs, supports most of the recent literature. Our finding that surgical resection does not impact survival in *IDH*wt LGGs is not broadly supported in the literature and requires further study. We speculate that disparate results regarding *IDH*wt LGGs may relate to the known clinical and molecular heterogeneity of this subgroup, especially among grade II *IDH*wt LGGs,^{8–10} and further stratification of this subgroup may be forthcoming in future WHO classification schemes.¹¹

Our study has limitations. Its retrospective design may have introduced unanticipated biases that affected the results. Our analysis might be limited by our cohort size and follow-up time, given the number of mortality events. For example, our relatively smaller cohort of *IDH*wt LGGs might have impacted our ability to discern survival benefits from surgical resection in this subgroup. Moreover, although we adhered to the latest WHO criteria, we did not include more detailed molecular testing of *IDH*wt LGGs (eg, *TERT* promoter mutation, chromosome 7/10 alterations, *EGFR* gene amplification) that might have allowed relevant substratification of this subgroup. Our volume measurements were performed using FLAIR and T2WI, and we recognize that edema and gliosis might have been inadvertently included within measurements due to their overlapping appearance with infiltrative gliomas on these imaging sequences. Finally, we measured tumor volumes using MR imaging scans acquired on various scanners with variable image quality; such variables might have affected the volume measurements.

CONCLUSIONS

The extent of surgical resection differentially impacts overall survival among patients with LGGs based on molecular subtypes as defined by the WHO. Increased extent of surgical resection significantly benefits patients with *IDH*-mutant LGGs, with the impact strongest for *IDH*mut-Noncodel astrocytomas. We found no survival benefit from greater surgical resection in *IDH*wt LGGs. Further study of the *IDH*wt subgroup, with larger cohort sizes and molecular substratification, is warranted.

Disclosures: Sohil H. Patel—RELATED: Grant: Radiological Society of North American Research Scholar Grant (RSCHI819)*; UNRELATED: Expert Testimony: LeClairRyan Law Firm.* M. Beatriz Lopes—UNRELATED: Royalties: Elsevier, Comments: book chapter royalties; Travel/Accommodations/Meeting Expenses Unrelated to Activities Listed: International Society of Neuropathology, Comments: travel expenses to attend the international meeting in Tokyo, Japan, October 2018. Camilo E. Fadul—UNRELATED: Grants/Grants Pending: Novartis, Comments: clinical trial for recurrent glioblastoma.* *Money paid to institution.

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